Disordered Systems and Random Graphs 1

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Overview

Lecture 1: introduction

- random graphs and phase transitionss
- the cavity method
- first/second moment method
- Belief Propagation and density evolution

Overview

Lecture 2: random 2-SAT

- the contraction method
- spatial mixing
- ► the Aizenman-Sims-Starr scheme
- the interpolation method

Overview

Lecture 3: group testing

- basics of Bayesian inference
- analysis of combinatorial algorithms
- spatial coupling
- information-theoretic lower bounds

Disordered systems



From glasses to random graphs [MP00]

- (spin) glasses are disordered materials rather than crystals
- lattice models are difficult to grasp even non-rigorously
- classical mean-field models: complete interaction
- diluted mean-field models: sparse random graph topology

Disordered systems



The binomial random graph $\mathbb{G} = \mathbb{G}(n, p)$ [ER60]

- vertex set x_1, \ldots, x_n
- connect any two vertices w/ probability $p = \frac{d}{n}$ independently
- ▶ local structure converges to Po(*d*) Galton-Watson tree

The Potts antiferromagnet



Definition

- fix d > 0, $q \ge 2$ and $\beta > 0$
- the Boltzmann distribution reads

$$\mu_{\mathbb{G},\beta}(\sigma) = \frac{1}{Z(\mathbb{G},\beta)} \prod_{vw \in E(\mathbb{G})} \exp(-\beta \mathbf{1}\{\sigma_v = \sigma_w\}) \quad (\sigma \in \{1,\dots,q\}^n)$$
$$Z(\mathbb{G},\beta) = \sum_{\tau \in \{1,\dots,q\}^n} \prod_{vw \in E(\mathbb{G})} \exp(-\beta \mathbf{1}\{\tau_v = \tau_w\})$$

The physics story: replica symmetry breaking



Replica symmetry

[KMRTSZ07]

- fix a large d and increase β
- for small β there are no extensive long-range correlations

$$\mu_{\mathbb{G},\beta}(\{\boldsymbol{\sigma}_{x_1} = \tau_1, \boldsymbol{\sigma}_{x_2} = \tau_2\}) \sim q^{-2} \qquad (\tau_1, \tau_2 \in \{1, \dots, q\})$$

▶ in fact, there is non-reconstruction and rapid mixing

The physics story: replica symmetry breaking



Dynamic replica symmetry breaking [KMRTSZ07]

still no extensive long-range correlations for moderate β

$$\mu_{\mathbb{G},\beta}(\{\boldsymbol{\sigma}_{x_1} = \tau_1, \boldsymbol{\sigma}_{x_2} = \tau_2\}) \sim q^{-2} \qquad (\tau_1, \tau_2 \in \{1, \dots, q\})$$

but there is reconstruction and torpid mixing

The physics story: replica symmetry breaking



Static replica symmetry breaking

[KMRTSZ07]

• for large β long-range correlations emerge

$$\mu_{\mathbb{G},\beta}(\{\boldsymbol{\sigma}_{x_1} = \tau_1, \boldsymbol{\sigma}_{x_2} = \tau_2\}) \neq q^{-2} \qquad (\tau_1, \tau_2 \in \{1, \dots, q\})$$

a few pure states dominate

The stochastic block model



The Potts model as an inference problem

[DKMZ11]

- choose a random colouring $\sigma^* \in \{1, ..., q\}^n$
- ▶ then choose a random graph G^{*} with

 $\mathbb{P}\left[\mathbb{G}^* = G \mid |E(\mathbb{G}^*)| = |E(G)|\right] \propto \mu_{G,\beta}(\boldsymbol{\sigma}^*)$

• given \mathbb{G}^* can we (partly) infer σ^* ?

Rigorous work

Techniques

- Classical random graphs techniques
 - method of moments
 - branching processes
 - large deviations
- Mathematial physics techniques
 - coupling arguments
 - exchangeable arrays and the cut metric
 - Belief Propagation and the contraction method
 - the interpolation method

Rigorous work



[ACO08,M12]

[G63,KRU13]

Success stories

- solution space geometry
- ► random *k*-SAT [AM02,AP03,COP16,DSS15]
- Iow-density parity check codes
- ► stochastic block model [AS15,M14,MNS13,MNS14,COKPZ16]
- group testing

[MTT08,COGHKL20]

Rigorous work



Theorem Let $\Lambda(x) = x \log x$ and

[COKPZ17]

$$\mathscr{B}_{q,\beta}^{*}(d) = \sup_{\pi} \mathscr{B}_{q,\beta,d}(\pi) \quad \text{where} \\ \mathscr{B}_{q,\beta,d}(\pi) = \mathbb{E}\bigg[\frac{\Lambda(\sum_{\sigma=1}^{q} \prod_{i=1}^{\gamma} 1 - (1 - e^{-\beta})\boldsymbol{\mu}_{i}^{(\pi)}(\sigma))}{q(1 - (1 - e^{-\beta})/q)^{\gamma}} - \frac{d}{2} \frac{\Lambda(1 - (1 - e^{-\beta})\sum_{\sigma=1}^{q} \boldsymbol{\mu}_{1}^{(\pi)}(\sigma)\boldsymbol{\mu}_{2}^{(\pi)}(\sigma))}{1 - (1 - e^{-\beta})/q}\bigg]$$

Then

$$d_{\text{cond}}(q,\beta) = \inf \left\{ d > 0 : \mathscr{B}_{q,\beta}^*(d) = \ln q + \frac{d}{2} \ln(1 - (1 - e^{-\beta})/q) \right\}.$$

The 2-SAT problem

- Boolean variables x_1, \ldots, x_n
- truth values +1 and -1
- four types of clauses:

 $x_i \lor x_j$ $x_i \lor \neg x_j$ $\neg x_i \lor x_j$ $\neg x_i \lor \neg x_j$

- ► a 2-SAT formula is a conjunction $\Phi = \bigwedge_{i=1}^{m} a_i$ of clauses
- $S(\Phi) = \text{set of satisfying assignments}$
- $\blacktriangleright Z(\Phi) = |S(\Phi)|$



Example

- $\bullet \quad \Phi = (\neg x_1 \lor x_2) \land (x_1 \lor x_3) \land (\neg x_2 \lor \neg x_3)$
- $Z(\Phi) = 2$ and $S(\Phi)$ consists of the two assignments

$\sigma_{x_1} = +1$	$\sigma_{x_2} = +1$	$\sigma_{x_3} = -1$
$\sigma_{x_1} = -1$	$\sigma_{x_2} = -1$	$\sigma_{x_3} = +1$





Computational complexity

2-SAT admits an efficient decision algorithm	[K67]
in fact, WalkSAT solves the problem efficiently	[P91]
the problem is NL-complete	[IS87,P94]
however, computing $\log Z(\Phi)$ is #P-hard	[V79]



Random 2-SAT

- for a fixed $0 < d < \infty$ let m = Po(dn/2)
- Φ = conjunction of *m* independent random clauses
- variable degrees have distribution Po(d)
- *Key questions:* is $Z(\Phi) > 0$ and if so, what is

$$\lim_{n \to \infty} \frac{1}{n} \log Z(\mathbf{\Phi}) \quad \Im$$

Prior work

- the threshold for $S(\Phi) = \emptyset$ occurs at d = 2 [CR92,G96]
- computation of $\log Z(\Phi)$ via replica/cavity method [MZ96]
- ► the scaling window [BBCKW01]
- partial results on 'soft' version

[T01,MS07,P14] [AM14]

• existence of a function $\phi(d)$ such that

$$\lim_{n \to \infty} \frac{\log Z(\mathbf{\Phi})}{n} = \phi(d)$$

for almost all $d \in (0, 2)$

The satisfiability threshold



Bicycles

• the clause $l \lor l'$ is logically equivalent to the two implications

$$l \lor l' \equiv (\neg l \to l') \land (\neg l' \to l)$$

• Φ is satisfiable unless there is an implication chain

$$x_i \rightarrow \cdots \rightarrow \neg x_i \rightarrow \cdots \rightarrow x_i$$

such chains are called *bicycles*

The satisfiability threshold



Theorem

[CR92,G96]

- If d < 2 then Φ does not contain a bicycle w.h.p.
- If d > 2 then Φ contains a bicycle w.h.p.

A naive attempt

- we aim to compute $\log Z(\Phi)$ for a typical Φ
- Jensen's inequality shows that

 $\log Z(\Phi) \le \log \mathbb{E}[Z(\Phi) \mid m] + o(n) \qquad \text{w.h.p.}$

The first moment

• computing $E[Z(\Phi) | m]$ is a cinch:

$$\mathrm{E}[Z(\mathbf{\Phi}) \mid \mathbf{m}] = 2^n \cdot \left(\frac{3}{4}\right)^m$$



$$\frac{1}{n}\log Z(\Phi) \le (1-d)\log 2 + \frac{d}{2}\log 3 \qquad \text{w.h.p}$$

The second moment

- this bound is tight if $E[Z(\mathbf{\Phi})^2] = O(E[Z(\mathbf{\Phi})]^2)$
- we calculate

$$E[Z(\mathbf{\Phi})^2 \mid \mathbf{m}] = \sum_{\sigma, \tau \in \{\pm 1\}^n} P[\mathbf{\Phi} \models \sigma, \mathbf{\Phi} \models \tau \mid \mathbf{m}]$$
$$= \sum_{\ell=-n}^n \sum_{\sigma, \tau: \sigma \cdot \tau = \ell} \left(\frac{1}{2} + \frac{(1+\ell/n)^2}{16}\right)^{\mathbf{m}}$$
$$= \sum_{\ell=-n}^n \binom{n}{(n+\ell)/2} \left(\frac{1}{2} + \frac{(1+\ell/n)^2}{16}\right)^{\mathbf{m}}$$



The second moment

► hence,

$$\frac{1}{n}\log \mathbb{E}[Z(\Phi)^2 \mid m] \sim \max_{-1 \le \alpha \le 1} H((1+\alpha)/2) + \frac{d}{2}\log\left(\frac{1}{2} + \frac{(1+\alpha)^2}{16}\right)$$

• at $\alpha = 0$ the above function evaluates to

$$\log 2 + d\log \frac{3}{4} \sim \frac{2}{n}\log \mathbb{E}[Z(\boldsymbol{\Phi}) \mid \boldsymbol{m}]$$

• therefore, we succeed iff the max is attained at $\alpha = 0$



The factor graph

- vertices x_1, \ldots, x_n represent variables
- vertices a_1, \ldots, a_m represent clauses
- the graph $G(\Phi)$ contains few short cycles
- ▶ locally $G(\Phi)$ resembles a Galton-Watson branching process



The Boltzmann distribution

• assuming $S(\mathbf{\Phi}) \neq \emptyset$ define

$$\mu_{\Phi}(\sigma) = \frac{\mathbf{1}\{\sigma \in S(\Phi)\}}{Z(\Phi)} \qquad (\sigma \in \{\pm 1\}^{\{x_1, \dots, x_n\}})$$

• let $\boldsymbol{\sigma} = \boldsymbol{\sigma}_{\Phi}$ be a sample from μ_{Φ}



Belief Propagation

define the variable-to-clause messages by

$$\mu_{\Phi, x \to a}(\sigma) = \mu_{\Phi-a}(\sigma_x = \sigma) \qquad (\sigma = \pm 1)$$





Belief Propagation

define the clause-to-variable messages by

$$\mu_{\Phi,a\to x}(\sigma) = \mu_{\Phi-(\partial x \setminus a)}(\sigma_x = \sigma) \qquad (\sigma = \pm 1)$$





The replica symmetric ansatz

The messages (approximately) satisfy

$$\mu_{\Phi,x \to a}(\sigma) \propto \prod_{b \in \partial x \setminus a} \mu_{\Phi,b \to x}(\sigma)$$

$$\mu_{\Phi,a \to x}(\sigma) \propto 1 - \mathbf{1} \left\{ \sigma \neq \operatorname{sign}(x,a) \right\} \mu_{\Phi,\partial a \setminus x}(-\operatorname{sign}(\partial a \setminus x))$$



The Bethe free entropy

► we expect that

$$\log Z(\mathbf{\Phi}) \sim \sum_{i=1}^{n} \log \sum_{\sigma=\pm 1} \prod_{a \in \partial x_i} \mu_{\mathbf{\Phi}, a \to x}(\sigma) + \sum_{i=1}^{m} \log \left(1 - \prod_{x \in \partial a_i} \mu_{\mathbf{\Phi}, x \to a_i}(-\operatorname{sign}(x, a_i)) \right) - \sum_{i=1}^{n} \sum_{a \in \partial x_i} \log \sum_{\sigma=\pm 1} \mu_{\mathbf{\Phi}, x \to a_i}(\sigma) \mu_{\mathbf{\Phi}, a_i \to x}(\sigma)$$



Density evolution

consider the empirical distribution of the messages:

$$\pi_{\mathbf{\Phi}} = \frac{1}{2m} \sum_{i=1}^{n} \sum_{a \in \partial x_i} \delta_{\mu_{\mathbf{\Phi}, x \to a}(+1)}$$

• d^+ , $d^- \sim \text{Po}(d/2)$, $\mu_0, \mu_1, \mu_2, \dots$ samples from π_{Φ}

$$\boldsymbol{\mu}_{0} \stackrel{\mathrm{d}}{=} \frac{\prod_{i=1}^{d_{+}} \boldsymbol{\mu}_{i}}{\prod_{i=1}^{d_{+}} \boldsymbol{\mu}_{i} + \prod_{i=1}^{d_{-}} \boldsymbol{\mu}_{i+d^{+}}}$$





Summary: the replica symmetric prediction[MZ96]For d < 2 there is a unique distribution π_d on (0, 1) s.t.

$$\boldsymbol{\mu}_{0} \stackrel{\text{d}}{=} \frac{\prod_{i=1}^{d_{+}} \boldsymbol{\mu}_{i}}{\prod_{i=1}^{d_{+}} \boldsymbol{\mu}_{i} + \prod_{i=1}^{d_{-}} \boldsymbol{\mu}_{i+d^{+}}}$$

and $\lim_{n\to\infty} n^{-1}\log Z(\Phi) = \mathscr{B}_d$ where

$$\mathscr{B}_d = \mathbf{E}\left[\log\left(\prod_{i=1}^{d_+} \boldsymbol{\mu}_i + \prod_{i=1}^{d_-} \boldsymbol{\mu}_{i+d_+}\right) - \frac{d}{2}\log(1 - \boldsymbol{\mu}_1 \boldsymbol{\mu}_2)\right]$$





Theorem

[ACOHKLMPZ20]

For *d* < 2 there is a unique distribution π_d on (0, 1) s.t.

$$\boldsymbol{\mu}_0 \stackrel{\mathrm{d}}{=} \frac{\prod_{i=1}^{d_+} \boldsymbol{\mu}_i}{\prod_{i=1}^{d_+} \boldsymbol{\mu}_i + \prod_{i=1}^{d_-} \boldsymbol{\mu}_{i+d^+}}$$

and $\lim_{n\to\infty} n^{-1}\log Z(\Phi) = \mathcal{B}_d$ where

$$\mathscr{B}_d = \mathrm{E}\left[\log\left(\prod_{i=1}^{\boldsymbol{d}_+} \boldsymbol{\mu}_i + \prod_{i=1}^{\boldsymbol{d}_-} \boldsymbol{\mu}_{i+\boldsymbol{d}_+}\right) - \frac{d}{2}\log(1 - \boldsymbol{\mu}_1 \boldsymbol{\mu}_2)\right]$$

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