

Wednesday, March 31

(start: Devang's presentation on the
Class Number Formula)

Toy example in probability:

Let X_1, X_2, X_3 be independent random variables,
uniform on $[-50, 50]$, $[-6, 6]$, $[-6, 6]$.

• Note that $P_i(X_i > X_j) = \frac{1}{2}$ for every
 $i \neq j$, $i, j \in \{1, 2, 3\}$.

• Three-way orderings:

• Suppose $X_1 > 6$. (44%)

$X_1 > X_2 > X_3$ 22%

$X_1 > X_3 > X_2$ 22%

• Suppose $X_1 < -6$ (44%)

$X_2 > X_3 > X_1$ 22%

$X_3 > X_2 > X_1$ 22%

• If we condition on $|X_1| \leq 6$: (12%)

- each of the six orderings is
equally likely by symmetry.

So overall, the three-way orderings are:

$X_1 > X_2 > X_3$ 24%

$X_1 > X_3 > X_2$ 24%

$X_2 > X_3 > X_1$ 24%

$X_3 > X_2 > X_1$ 24%

$X_2 > X_1 > X_3$ 2%

$X_3 > X_1 > X_2$ 2%

Next example: X_1, X_2, X_3 independent normal
random variables with mean 0,
variances $\alpha^2, \beta^2, \gamma^2$.

• By symmetry:

$$P_i(X_i > X_j > X_k) = P_i(X_k > X_j > X_i).$$

• Turns out one can calculate

$P_i(X_1 > X_2 > X_3)$ exactly.

Next example: X_1, X_2, X_3 independent normal random variables with mean 0, variances $\alpha^2, \beta^2, \gamma^2$.

• Density function of X_1 is $\frac{1}{\alpha\sqrt{2\pi}} e^{-x^2/2\alpha^2}$ (etc.)

We want to calculate $\Pr(X_1 < X_2 < X_3) =$

$$\int_{-\infty}^{\infty} \int_x^{\infty} \int_y^{\infty} \frac{1}{\alpha\beta\gamma(2\pi)^{3/2}} e^{-\frac{x^2}{2\alpha^2} - \frac{y^2}{2\beta^2} - \frac{z^2}{2\gamma^2}} dz dy dx.$$

Change of variables $y = x + s, z = x + s + t$:

$$\int_{-\infty}^{\infty} \int_0^{\infty} \int_0^{\infty} \frac{1}{\alpha\beta\gamma(2\pi)^{3/2}} e^{-\frac{x^2}{2\alpha^2} - \frac{(x+s)^2}{2\beta^2} - \frac{(x+s+t)^2}{2\gamma^2}} dt ds dx.$$

Fact: $\int_{-\infty}^{\infty} e^{-ax^2 - bx} dx = \int_{-\infty}^{\infty} e^{-a(x - \frac{b}{2a})^2 + \frac{b^2}{4a}} dx$

$$= e^{\frac{b^2}{4a}} \int_{-\infty}^{\infty} e^{-a(x - \frac{b}{2a})^2} dx = e^{\frac{b^2}{4a}} \sqrt{\frac{\pi}{a}}.$$

Thus

$$\frac{1}{\alpha\beta\gamma(2\pi)^{3/2}} \int_{-\infty}^{\infty} e^{-\frac{x^2}{2\alpha^2} - \frac{(x+s)^2}{2\beta^2} - \frac{(x+s+t)^2}{2\gamma^2}} dx = \frac{1}{2\pi\tau} \exp\left(-\frac{s^2(\beta^2 + \gamma^2) - 2st(\alpha^2 + \gamma^2) + t^2(\alpha^2 + \beta^2)}{2\tau^2}\right),$$

where $\tau^2 = \alpha^2\beta^2 + \beta^2\gamma^2 + \gamma^2\alpha^2$.

• Remaining integral is $\int_0^{\infty} \int_0^{\infty} dt ds$,

which is an "orthant probability":

$\Pr(\text{some 2nd normal variable lies in } [0, \infty])$

• possible to diagonalize the quadratic form to $e^{-(w^2 + u^2)}$ by linear change

of variables; region of integration

turns into a sector in \mathbb{R}^2 .

trivial to integrate in polar coordinates.

Final answer:

$$\Pr(X_1 < X_2 < X_3) = \frac{1}{2\pi} \arctan \frac{\tau}{\beta^2}.$$

$$\Pr(X_1 < X_2 < X_3) = \frac{1}{2\pi} \arctan \frac{\sqrt{\alpha^2 \beta^2 + \beta^2 \gamma^2 + \gamma^2 \alpha^2}}{\beta^2} \quad (*)$$

If $\alpha^2, \beta^2, \gamma^2$ are all close to each other:

$$\text{write } \alpha^2 = \sigma^2(1+\varepsilon), \beta^2 = \sigma^2(1+\delta),$$

$$\gamma^2 = \sigma^2(1+\eta). \text{ Linear approximation to } (*)$$

* $(\varepsilon, \delta, \eta) = (0, 0, 0)$ gives

$$\Pr(X_1 < X_2 < X_3) = \frac{1}{6} + \frac{\varepsilon - 2\delta + \eta}{8\pi\sqrt{3}} + O(\varepsilon^2 + \delta^2 + \eta^2)$$

• Literature contains closed forms for 3- and 4-dimensional orthant probabilities of multivariate normal random variables — non-diagonal ok (X_1, \dots, X_4 don't have to be independent).

— When we estimate $\int_{q_1, a_1, \dots, a_r}$ by $A(X_1 > \dots > X_r)$, we can set roughly

$\int_{q_1, a_1, \dots, a_r}$ = (orthant probability in closed form) $\rightarrow O\left(\frac{1}{q}\right)$.

Example Lamzouri, "Prime number races with three or more competitors"

Corollary 2.3:

$$\int_{q_1, a_1, a_2, a_3} = \frac{1}{6} + \frac{1}{4\sqrt{\pi}} \frac{c(q, a_3) - c(q, a_1)}{\sqrt{V(q)}} + \frac{1}{4\pi\sqrt{3}} \frac{B_2(a_1, a_2) - 2B_2(a_1, a_3) + B_2(a_2, a_3)}{\sqrt{V(q)}} + O(b^2)$$

$$c(q, a) = -1 + \#\{b^2 \equiv a \pmod{q}\}$$

$$V(q) = 2 \sum_{x \pmod{q}} \sum_{\substack{y > 0 \\ U(x+y, x) = 0}} \frac{1}{\frac{1}{4} + y^2} \sim 2\phi(q) \log q$$

$$B_2(a, b) = 2 \sum_{x \pmod{q}} (X(bx^{-1}) + X(bx)) \sum_{r > 0} \frac{1}{\frac{1}{4} + r^2}$$