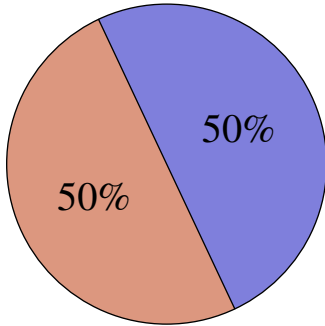


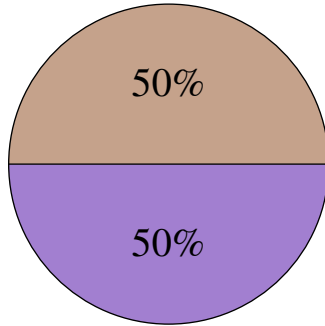
The three-way (mod 12) race among 5, 7, and 11:

$$\pi(x; 12, 5) > \pi(x; 12, 7)$$



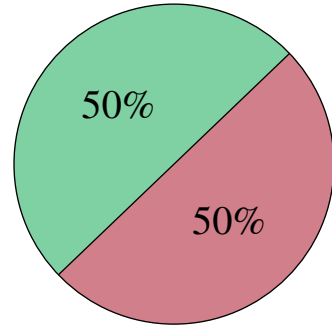
$$\pi(x; 12, 7) > \pi(x; 12, 5)$$

$$\pi(x; 12, 5) > \pi(x; 12, 11)$$



$$\pi(x; 12, 11) > \pi(x; 12, 5)$$

$$\pi(x; 12, 7) > \pi(x; 12, 11)$$

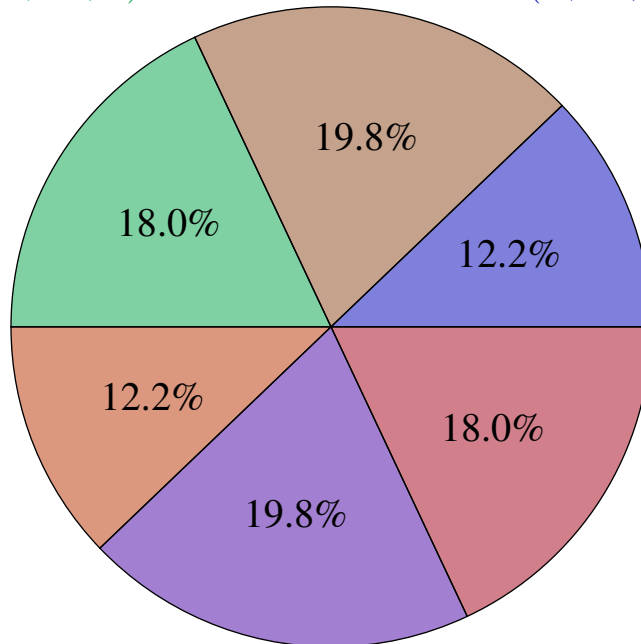


$$\pi(x; 12, 11) > \pi(x; 12, 7)$$

$$\pi(x; 12, 5) > \pi(x; 12, 7) > \pi(x; 12, 11)$$

$$\pi(x; 12, 7) > \pi(x; 12, 5) > \pi(x; 12, 11)$$

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$$\pi(x; 12, 7) > \pi(x; 12, 11) > \pi(x; 12, 5)$$

$$\pi(x; 12, 11) > \pi(x; 12, 5) > \pi(x; 12, 7)$$

$$\pi(x; 12, 11) > \pi(x; 12, 7) > \pi(x; 12, 5)$$

Friday, March 17

Toy example in probability:

Let X_1, X_2, X_3 be independent random variable, uniform on $[-50, 50]$, $[0, 6]$, and $[0, 6]$.

• Note that $\Pr(X_i > X_j) = \frac{1}{2}$ for any $i \neq j$, $(i, j) \in \{1, 2, 3\}$.

• Three-way orderings:

• Suppose $X_1 > 6$ (44%).

$X_1 > X_2 > X_3$ (22%)

$X_1 > X_3 > X_2$ (22%)

• Suppose $X_1 < -6$ (44%)

$X_2 > X_3 > X_1$ (22%)

$X_3 > X_2 > X_1$ (22%)

• Suppose $X_1 \in [0, 6]$.

– each of the 3! three-way orderings is equally likely: 2%.

∴ overall, the three-way ones gives:

$X_1 > X_2 > X_3$: 24%

$X_1 > X_3 > X_2$: 24%

$X_2 > X_3 > X_1$: 24%

$X_3 > X_2 > X_1$: 24%

$X_2 > X_1 > X_3$: 2%

$X_3 > X_1 > X_2$: 2%

Next example: X_1, X_2, X_3 independent normal random variables with mean 0, variances $\alpha^2, \beta^2, \gamma^2$.

– Each X_j vs. X_i is 50/50 race.

• Symmetry: $\Pr(X_i > X_j > X_k) = \Pr(X_k > X_j > X_i)$.

Turns out we can calculate $\Pr(X_1 < X_2 < X_3)$.

• Density function of X_i is $\frac{1}{\alpha\sqrt{2\pi}} e^{-x^2/2\alpha^2}$ (etc.)

We want to calculate $\Pr(X_1 < X_2 < X_3) =$

$$\int_{-\infty}^{\infty} \int_x^{\infty} \int_y^{\infty} \frac{1}{\alpha\beta\gamma(2\pi)^{3/2}} e^{-\frac{x^2}{2\alpha^2} - \frac{y^2}{2\beta^2} - \frac{z^2}{2\gamma^2}} dz dy dx$$

Change variables $y = x + s, z = x + s + t$:

$$\int_{-\infty}^{\infty} \int_0^{\infty} \int_0^{\infty} \frac{1}{\alpha\beta\gamma(2\pi)^{3/2}} e^{-\frac{x^2}{2\alpha^2} - \frac{(x+s)^2}{2\beta^2} - \frac{(x+s+t)^2}{2\gamma^2}} dt ds dx$$

Fact: $\int_{-\infty}^{\infty} e^{-ax^2 - bx} dx = \int_{-\infty}^{\infty} e^{-a(x + \frac{b}{2a})^2 + \frac{b^2}{4a}} dx$

$$= e^{\frac{b^2}{4a}} \int_{-\infty}^{\infty} e^{-a(x + \frac{b}{2a})^2} dx = e^{\frac{b^2}{4a}} \sqrt{\frac{\pi}{a}}$$

Thus

$$\frac{1}{\alpha\beta\gamma(2\pi)^{3/2}} \int_{-\infty}^{\infty} e^{-\frac{x^2}{2\alpha^2} - \frac{(x+s)^2}{2\beta^2} - \frac{(x+s+t)^2}{2\gamma^2}} dx =$$

$$= \frac{1}{2\pi\tau} \exp\left(-\frac{s^2(\beta^2 + \gamma^2) - 2st(\alpha^2 + \gamma^2) + t^2(\alpha^2 + \beta^2)}{2\tau^2}\right),$$

where $\tau^2 = \alpha^2\beta^2 + \beta^2\gamma^2 + \gamma^2\alpha^2$.

• Remaining integral is $\int_0^{\infty} \int_0^{\infty} dz dy$,

which is an "orthant probability" —

$\Pr(\text{some } n \text{ normal random variable lies in } [0, \infty)^2)$.

— Possible to diagonalize the quadratic form $(*)$ to $e^{-(v^2 + w^2)}$ by linear

change of variables; region of integration becomes a sector in \mathbb{R}^2 .

— trivial to integrate in polar coords.

Final answer:

$$\Pr(X_1 < X_2 < X_3) = \frac{1}{2\pi} \arctan \frac{\tau}{\beta^2}.$$

$$\Pr(X_1 < X_2 < X_3) = \frac{1}{2\pi} \arctan\left(\frac{\sqrt{\alpha^2\beta^2 + \beta^2\gamma^2 + \gamma^2\alpha^2}}{\beta^2}\right) \quad (*)$$

If $\alpha^2, \beta^2, \gamma^2$ are all close to each other:
 $\alpha^2 = \sigma^2(1+\epsilon), \beta^2 = \sigma^2(1+\delta), \gamma^2 = \sigma^2(1+\eta)$.

Linear approximation to (*) at $(\epsilon, \delta, \eta) = (0, 0, 0)$ is

$$\Pr(X_1 < X_2 < X_3) = \frac{1}{6} + \frac{\epsilon - 2\delta + \eta}{8\pi\sqrt{3}} + O(\epsilon^2 + \delta^2 + \eta^2)$$

• Literature contains closed formulas for 3- and 4-dimensional arthant probabilities of multivariate normal random variable with mean 0 — or if nondiagonal (X_1, \dots, X_4 don't have to be independent)

— When we estimate $\delta_{q; a_1, \dots, a_r}$ = log-likelihood of $\{x: \pi(x; q, a_1) > \dots > \pi(x; q, a_r)\}$

by $\Pr(X_1 > \dots > X_r)$, we get roughly $\delta_{q; a_1, \dots, a_r} = (\text{some arthant probability in } \mathbb{R}^r) + O(1/q)$.

Example: Lamzouri, "Prime number races w/ 3 or more competitors", Corollary 2.3:

$$\delta_{q; a_1, a_2, a_3} = \frac{1}{6} + \frac{1}{4\sqrt{\pi}} \frac{c(q, a_3) - c(q, a_1)}{\sqrt{V(q)}} + \frac{1}{4\pi\sqrt{3}} \frac{B_q(a_1, a_2) - 2B_q(a_1, a_3) + B_q(a_2, a_3)}{V(q)} + O(b^2/n^2)$$

$$c(q, a) = -1 + \#\{b^2 \equiv a \pmod{q}\}$$

$$V(q) = 2 \sum_{\substack{x \pmod{q} \\ (\frac{x}{q} + ir, x) = 0}} \sum_{r > 0} \frac{1}{4+r^2} \sim 2\pi^2/q \log q$$

$$B_q(a, b) = 2 \sum_{x \pmod{q}} (\chi(\bar{a}^{-1}) + \chi(\bar{b}^{-1})) \sum_{r > 0} \frac{1}{4+r^2}$$